### 1.3 Review of epidemiological tables

RR, OR, confounding, stratification, trend

## Risk

The simplest measure of risk (probability) is estimated from the proportion that is observed

## Example:

If in a cohort of 11034 people, there were 189 who had an MI Risk of $\mathrm{MI}=189 / 11034=.0171=1.7 \%$

Can calculate risk from cohort or cross-sectional study but not from case-control study (traditional wisdom!)

## Comparing Proportions (Chi-square test)

For any size table, the Chi-square test answers the question: "is there an association between the two factors

How?

By calculating the "expected table" if no association and comparing to actual observed table

If the "discrepency" is large, conclude there is an association

How large? depends on size of table

## Example of two-by-two table

| Graft Rejection | Marrow Cell Dose(108 cells/kg) |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | $<3.0$ | X3.0 | Total |
|  | Yes | 17 | 4 | 21 |
|  | No | 19 | 28 | 47 |
|  | Total |  |  | 68 |
|  |  |  | Colum | Grand |
|  |  |  | Totals | Total |

Question: Is there evidence that marrow dose is associated with GVHD?
i.e. is the above Table what we would expect even
if there is no association?

## Overall GVHD rate=21/68

-so if no association, we would expect:
-21/68 of 36 in $<3.0$ group to get GVHD
-and $21 / 68$ of 32 in the other group
i.e. we expect


## Compare:

|  |  |  |  |
| :---: | :---: | :---: | :---: |
|  |  | Observed |  |
|  | $<3.0$ | X3.0 |  |
| Yes | 17 | 4 |  |
| No | 19 | 28 |  |
|  |  |  |  |

Expected (if no association)

| $<3.0$ | X3.0 |
| :---: | :---: |
| 11.1 | 9.9 |
| 24.9 | 22.1 |

$$
\begin{aligned}
\text { Discrepency } & =\sum \frac{(\text { observed -expected })^{2}}{\text { expected }} \\
& =\frac{(17-11.1)^{2}}{11.1}+\frac{(4-9.9)^{2}}{9.9}+\frac{(19-24.9)^{2}}{24.9}+\frac{(28-22.1)^{2}}{22.1} \\
& =9.64 \\
& \text { Is this large or small? }
\end{aligned}
$$

## Tests of discrepancy

If no association, discrepency has a Chi-Squared distribution with 1 degree of freedom, $\chi^{2}{ }_{(1)}$ (square of standard normal!)

our result has $p$-value $<.01$, so is unlikely to be simply due to randomness
we conclude that there is a difference in the two groups
This is called "Pearson's Chi square test"

For any size table ( R rows, C columns), can also construct an observed and expected table, but under null hypothesis, distribution of discrepency now $\chi^{2}(\mathrm{R}-1 \times \mathrm{C}-1)$

## For 2-by-2 table, can also compare proportions using Relative Risk (Risk Ratio)

N Eng. Jour. Med. 1988 (262-264) Physicians Health Study: randomised trial of regular use of aspirin and 5-year MI rate

|  |  | Yes | No |
| :--- | :---: | :---: | :---: | Total $10.11,034$

Risk of MI for placebo $=189 / 11034=.0171$
Risk of MI for aspirin $=104 / 11037=.0094$
$\Rightarrow \mathrm{RR}=1.82$

To assess significance: confidence interval (or Chi-square test)

## Can also compare Odds (Odds Ratio)

## Odds <br> "the ratio of successes to failures"

## Example: Dental Analgesic Trial

|  | Relief |  |
| :---: | :---: | :---: |
|  | Y | N |
| Active | 24 | 6 |
| Placebo | 3 | 17 |

Odds of relief in Active $=24 / 6$

Odds of relief in Placebo=3/17

Odds Ratio of relief in Active compared to Placebo :
$24 / 6 \div 3 / 17=22.7(p<.0001)!$
$\mathrm{OR}=1 \Rightarrow$ no association (like $\mathrm{RR}=1$ )
OR $>1 \Rightarrow+$ association treatment \& relief
$\mathrm{OR}<1 \Rightarrow$ - association treatment and pain

## Reverse the Question

Compute OR of being on active treatment for those with pain relief compared to those with no pain relief

Example: Dental Analgesic Trial

|  | Relief |  |
| :---: | :---: | :---: |
|  | $Y$ | N |
| Active | 24 | 6 |
| Placebo | 3 | 17 |

## same!

i.e. we can calculate and interpret OR from case-control studies

## Traditional wisdom:

Only the OR is valid from case-control study

## Relationship between OR and RR

Text book wisdom: if disease is rare then $O R \cong R R$
(as in placebo vs. aspirin example)

## Crude and stratified OR

## (example from Zang and Wynder*)

|  | Overall |  |  | Smoker |  |  | Non-smoker |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Lung Cancer |  | 3006 | Lung Cancer |  |  | Lung Cancer |  |  |
|  | Yes | No |  | Yes | No |  | Yes | No |  |
| Heavy Drinker Light Drinker | 1057 | 1949 |  | 786 | 665 | 1451 | 271 | 1248 | 1555 |
|  | 1896 | 6220 | 8116 | 702 | 591 | 1293 | 1194 | 5629 | 6823 |
|  | 2953 | 8169 | 11122 | 1488 | 1256 | 2744 | 1463 | 6913 | 8378 |
|  | $\mathrm{OR}=1.78$ |  |  | $\mathrm{OR}=1.0$ |  |  | $\mathrm{OR}=1.0$ |  |  |

Illustration of confounding

[^0]
## Smoking is a potential "confounder": Eme tisi Karolinska Institutet



## Example for continuous variables

Where a "predictor" X is presumed to be associated causally with outcome, $Y$, but there is an additional variable, $Z$, that is associated with both $X$ and $Y$


Wainer et al. Giving the Finger to Dating Services. CHANCE, 21:3, 59-61.

## Examples of potential confounders

- Studied association
- Birth weight and adult heart disease
- Vitamin D and Mycardial infarction


## Confounder?

maternal smoking adult BMI /weight
fast food consumption sun exposure?

- Prenatal tobacco and own tobacco use Parental smoking in childhood, gestational age?


## Mantel-Haenszel OR (common OR)

 for binarv outcome and exposure|  | Stratum 1 |  |
| ---: | :---: | :---: |
|  | Disease |  |
| Exposed | No |  |
| Enexposed | $a_{1}$ | $b_{1}$ |
| Uny | $c_{1}$ | $d_{1}$ |
|  |  |  |


|  | Stratum 2 |  |
| :---: | :---: | :---: |
|  | Disease |  |
| Expo | No |  |
| Enexposed | $a_{2}$ | $b_{2}$ |
| Uny | $c_{2}$ | $d_{2}$ |
|  |  |  |

$n_{1}$

$$
\widehat{O R}_{M H}=\frac{\frac{a_{1} d_{1}}{n_{1}}+\frac{a_{2} d_{2}}{n_{2}}}{\frac{b_{1} c_{1}}{n_{1}}+\frac{b_{2} c_{2}}{n_{2}}} \longleftarrow \text { Main diagonals }
$$

$$
\begin{aligned}
\widehat{O R}_{M H} & =\frac{\frac{786 \times 591}{2744}+\frac{271 \times 5629}{8378}}{\frac{665 \times 702}{2744}+\frac{1248 \times 1194}{8378}} \\
& =351.367 / 347.988 \\
& =1.0097
\end{aligned}
$$

Can have any number of strata

## The MH-OR estimates the common OR

So before computing $\mathrm{OR}_{\mathrm{MH}}$, we need to test whether it is reasonable to assume a common OR
"Tests of homogeneity"

- All are similar in spirit to the simple Chi-squared test of association
- compare the observed data in each stratum to what would be expected if there was a common OR (i.e. the overall OR)
- compare the total "discrepancy" to Chi-square distribution

Provided in statistical calculators (Openepi.com) and software

## Example: Framingham data



Crude OR $=\mathbf{2 . 1 2}$

|  | Ages 42-51 |  |  |  | Ages 52-61 |  |  |  |  |  | Ages 62-71 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Diabetes |  |  |  |  |  |  | etes |  |  |  |  | tes |  |
|  |  | Yes | No |  |  |  | Yes | No |  |  |  | Yes | No |  |
| CHD | Yes | 1 | 32 | $\begin{gathered} 33 \\ 1581 \end{gathered}$ | CHD | Yes | 4 | 74 | $\begin{gathered} 78 \\ 1268 \end{gathered}$ | CHD | Yes | 7 | 69 | 76 |
|  | No | 28 | 1553 |  |  | No | 47 | 1221 |  |  | No | 29 | 440 | 469 |
|  |  | 29 | 1585 | 1614 |  |  | 51 | 1295 | 1346 |  |  | 36 | 509 | 545 |
|  | $\mathrm{OR}=1.73$ |  |  |  |  | $\mathrm{OR}=1.40$ |  |  |  |  | $\mathrm{OR}=1.54$ |  |  |  |

## Test of homogeneity $\chi^{2}=.04, p=.98$

Mantel Haenszel OR $=1.5$

## Example: Framingham data (cont...)



Crude OR $=2.12$

## Mantel Haenszel OR $=1.5$

Conclude:
Confounding by age (adjusted differs by $>10 \%$ from crude)

## Control of Confounding

Removing spurious associations from related variables can be done at the design stage, and/or the analysis stage.
confounding is due to "imbalance", so idea is to
"balance" the design

## Control of confounding at design stage

## Restriction

- Confounding cannot occur if the factor does not vary. For example if the study is limited to non-smoking women, then smoking and gender cannot be confounding variables.
- Restriction also limits the participants/interpretation of the study. Often partial restriction is used.


## Matching

Later lecture

## Control of confounding at analysis stage

- Stratification (as shown for age groups earlier)
- Calculate adjusted OR (Mantel-Haenszel)
- Use "logistic regression" (more later) in a statistical package


## Effect modification (also called interaction)

When the effect of exposure is different in different strata ( test of homogeneity provides evidence against a common OR),

We say:
the effect is "modified" by the stratum
there is an "interaction"

Now a "common" OR not meaningful!!

If only a few strata, report the OR in each

## Example: Framingham data

| Prev. Hypertension | Diabetes |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | Yes | No |  |
|  | Yes | 67 | 1237 | 1304 |
|  | No | 49 | 2152 | 2201 |
|  |  | 116 | 3389 | 3505 |

Crude OR $=2.38$

|  | ales |  |  |  |  | nales |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Dia | etes |  |  |  |  | etes |  |
|  |  | Yes | No | $\begin{aligned} & 545 \\ & 983 \end{aligned}$ |  |  | Yes | No |  |
| Prev. Hypertension | $\begin{aligned} & \text { Yes } \\ & \text { No } \end{aligned}$ | 25 | 520 |  | Prev. Hypertension | Yes | 42 | 717 | 759 |
|  |  | 32 | 951 |  |  | No | 17 | 1201 | 1218 |
|  |  | 57 | 1471 | 1528 |  |  | 59 | 1918 | 1977 |
| $\mathrm{OR}=1.43$ |  |  |  |  | $\mathrm{OR}=4.14$ |  |  |  |  |

Test of homogeneity $\chi^{2}=7.13, \mathrm{p}=.008$

## Dose-response: test of trend

When the exposure is more than two levels and categories are ordered (e.g. age groups), may be a steady increase/ decrease in the risk with the 'dose' of exposure.

Important evidence: one of the Bradford-Hill criteria for causation

Return to alcohol and lung cancer (and smoking)example:


## Dose-response: test of trend

| Lung Cancer |  |  |  |
| :---: | :---: | :---: | :---: |
|  | Yes | No | Odds |
| $<1$ | 1090 | 3976 | . 274 |
| 1-3.9 | 806 | 2244 | . 359 |
| 4-6.9 | 378 | 783 | . 482 |
| $\geq 7$ | 679 | 1166 | . 582 |
|  | 2953 | 8169 |  |



Test of (any) association

$$
\left.\chi^{2}=186.8, \text { p-value }<.000001 \text { (from } \chi^{2} \text { with } 3 \text { d.o.f. }\right)
$$

## $\chi^{2}$ for trend (generally more powerful)



Where $\mathrm{a}_{\mathrm{i}}=$ cases in each stratum
$\mathrm{x}_{\mathrm{i}}=$ scores in the strata
$\mathrm{N}=$ total number of subjects
(cases + controls)

A = total number of cases
$\mathrm{p}=$ overall proportion cases

Expression in [] in numerator=
Total score for cases - (no. of cases) (average score overall)
Expression in [] in denominator = variance of score (avg. of square - square of avg.)

Under Null Hypothesis (no trend) this has Chi-squared distribution with 1 degree of freedom

## Test for trend (cont.)

Note we need to use scores: common to use midpoints.

For equally spaced strata, $1,2,3 \ldots$ give the same result

If no natural scores, can simply use $1,2,3 \ldots$.

Chi-squared test not sensitive to the choice.

## What if the exposure is continuous? <br> (e.g. age, blood pressure, biomarker levels...)

## Summary

- The effect of a risk factor on disease risk is usually measured by comparing prevalence, incidence, cumulative incidence or odds
- Comparisons in risk are most often based on relative difference, so by comparing the risk/odds of disease among exposed with the risk/odds among unexposed, e.g. RR or OR
- When comparing proportions across groups Chi-square tests are often used as a first test
$\rightarrow$ However, only gives p-values and no measure of association


## Summary (cont.)

- We looked at association between binary outcome and a single binary explanatory variable of interest
- Then we considered one explanatory variable and a confounder or stratum variable
$\rightarrow$ Test of homogeneity
$\rightarrow$ adjusted/common OR where appropriate
- Dose-reponse (test for trend)
- In practice we are often interested in a number of explanatory variables (independent risk factors, confounders, effect modifiers). So, after examining one-by-one ("univariate" analysis), we need to model their joint effect:

Logistic regression (later lecture)


[^0]:    * Zang E and Wynder E. Preventive Med 3, 359-370,2001

